# Tutorial on Bandit Games 

Sébastien Bubeck

Microsoft ${ }^{*}$<br>Research<br>PRINCETON UNIVERSITY

## Online Learning with Full Information

Adversary


Player

## Online Learning with Full Information

Adversary


## Online Learning with Full Information

Adversary


## Online Learning with Full Information

Adversary


Player

## Online Learning with Full Information

Adversary

loss suffered: $\ell_{A}$

$$
A \in\{1, \ldots, d\}
$$

Player

## Online Learning with Full Information



Online Learning with Bandit Feedback

Adversary


Player

## Online Learning with Bandit Feedback

Adversary


Player

## Online Learning with Bandit Feedback

Adversary $\longrightarrow$


Player

## Online Learning with Bandit Feedback



Player

## Online Learning with Bandit Feedback


loss suffered: $\ell_{A}$

Player

## Online Learning with Bandit Feedback



Computer Go


Packets routing


Brain computer interface


Ads placement


Medical trials


Dynamic allocation


## A little bit of advertising



E
S. Bubeck and N. Cesa-Bianchi.

Regret Analysis of Stochastic and Nonstochastic Multi-armed Bandit Problems.
Foundations and Trends in Machine Learning, Vol 5: No 1, 1-122, 2012.

For each round $t=1,2, \ldots, n$;
(1) The player chooses an $\operatorname{arm} I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t}, t}$, and observes:

Goal: Minimize the cumulative loss incured. We consider the regret:


For each round $t=1,2, \ldots, n$;
(1) The player chooses an arm $I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(3) Simultaneously the adversary chooses a loss vector
(3) The player incurs the loss $\ell_{l_{t}, t}$, and observes:

Goal: Minimize the cumulative loss incured. We consider the regret:

For each round $t=1,2, \ldots, n$;
(1) The player chooses an $\operatorname{arm} I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t, t}}$, and observes:

Goal: Minimize the cumulative loss incured. We consider the regret:

For each round $t=1,2, \ldots, n$;
(1) The player chooses an $\operatorname{arm} I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t}, t}$, and observes:

- The loss vector $l_{t}$ in the full information setting.
- Only the loss incured $\ell_{I_{t}, t}$ in the bandit setting.

Goal: Minimize the cumulative loss incured. We consider the regret:

For each round $t=1,2, \ldots, n$;
(1) The player chooses an $\operatorname{arm} I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t}, t}$, and observes:

- The loss vector $\ell_{t}$ in the full information setting.

Goal: Minimize the cumulative loss incured. We consider the regret:

For each round $t=1,2, \ldots, n$;
(1) The player chooses an arm $I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t}, t}$, and observes:

- The loss vector $\ell_{t}$ in the full information setting.
- Only the loss incured $\ell_{I_{t}, t}$ in the bandit setting.

Goal: Minimize the cumulative loss incured. We consider the regret:

For each round $t=1,2, \ldots, n$;
(1) The player chooses an $\operatorname{arm} I_{t} \in\{1, \ldots, d\}$, possibly with the help of an external randomization.
(2) Simultaneously the adversary chooses a loss vector $\ell_{t}=\left(\ell_{1, t}, \ldots, \ell_{d, t}\right) \in[0,1]^{d}$.
(3) The player incurs the loss $\ell_{\ell_{t}, t}$, and observes:

- The loss vector $\ell_{t}$ in the full information setting.
- Only the loss incured $\ell_{I_{t}, t}$ in the bandit setting.

Goal: Minimize the cumulative loss incured. We consider the regret:

$$
R_{n}=\mathbb{E} \sum_{t=1}^{n} \ell_{\ell_{t}, t}-\min _{i=1, \ldots, d} \mathbb{E} \sum_{t=1}^{n} \ell_{i, t}
$$

## Exponential Weights (EW, EWA, MW, Hedge, ect)

Draw $I_{t}$ at random from $p_{t}$ where

$$
p_{t}(i)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \ell_{j, s}\right)}
$$

## Theorem (Cesa-Bianchi, Freund, Haussler, Helmbold, Schapire and Warmuth [1997])

## Exponential Weights (EW, EWA, MW, Hedge, ect)

Draw $I_{t}$ at random from $p_{t}$ where

$$
p_{t}(i)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \ell_{j, s}\right)}
$$

## Theorem (Cesa-Bianchi, Freund, Haussler, Helmbold, Schapire and Warmuth [1997])

Exp satisfies

$$
R_{n} \leq \sqrt{\frac{n \log d}{2}}
$$

## Moreover for any strategy,

## Exponential Weights (EW, EWA, MW, Hedge, ect)

Draw $I_{t}$ at random from $p_{t}$ where

$$
p_{t}(i)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \ell_{j, s}\right)}
$$

## Theorem (Cesa-Bianchi, Freund, Haussler, Helmbold, Schapire and Warmuth [1997])

Exp satisfies

$$
R_{n} \leq \sqrt{\frac{n \log d}{2}}
$$

Moreover for any strategy,

$$
\sup _{\text {adversaries }} R_{n} \geq \sqrt{\frac{n \log d}{2}}+o(\sqrt{n \log d}) .
$$

The one-slide-proof

$$
w_{t}(i)=\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right), \quad w_{t}=\sum_{i=1}^{d} w_{t}(i), \quad p_{t}(i)=\frac{w_{t}(i)}{W_{t}}
$$

The one-slide-proof

$$
\begin{aligned}
w_{t}(i)=\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right), W_{t} & =\sum_{i=1}^{d} w_{t}(i), \quad p_{t}(i)=\frac{w_{t}(i)}{W_{t}} \\
\log \frac{W_{n+1}}{W_{1}}=\log \left(\frac{1}{d} \sum_{i=1}^{d} w_{n+1}(i)\right) & \geq \log \left(\frac{1}{d} \max _{i} w_{n+1}(i)\right) \\
& =-\eta \min _{i} \sum_{t=1}^{n} \ell_{i, t}-\log d
\end{aligned}
$$

The one-slide-proof

$$
\begin{aligned}
& w_{t}(i)=\exp \left(-\eta \sum_{s=1}^{t-1} \ell_{i, s}\right), W_{t}=\sum_{i=1}^{d} w_{t}(i), \quad p_{t}(i)=\frac{w_{t}(i)}{W_{t}} \\
& \log \frac{W_{n+1}}{W_{1}}=\log \left(\frac{1}{d} \sum_{i=1}^{d} w_{n+1}(i)\right) \geq \log \left(\frac{1}{d} \max _{i} w_{n+1}(i)\right) \\
&=-\eta \min _{i} \sum_{t=1}^{n} \ell_{i, t}-\log d \\
& \log \frac{W_{n+1}}{W_{1}}=\sum_{t=1}^{n} \log \frac{W_{t+1}}{W_{t}}=\sum_{t=1}^{n} \log \left(\sum_{i=1}^{d} \frac{w_{t}(i)}{W_{t}} \exp \left(-\eta \ell_{i, t}\right)\right) \\
&=\sum_{t=1}^{n} \log \left(\mathbb{E} \exp \left(-\eta \ell_{l_{t}, t}\right)\right) \\
& \leq \sum_{t=1}^{n}\left(-\eta \mathbb{E} \ell_{l_{t}, t}+\frac{\eta^{2}}{8}\right)
\end{aligned}
$$

Magic trick for bandit feedback

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{p_{t}(i)} \mathbb{1}_{I_{t}=i}
$$

is an unbiased estimate of $\ell_{i, t}$. We call Exp3 the Exp strategy run on the estimated losses.

Theorem (Auer, Cesa-Bianchi, Freund and Schapire 2003]
Exp3 satisfies:

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{p_{t}(i)} \mathbb{1}_{I_{t}=i}
$$

is an unbiased estimate of $\ell_{i, t}$. We call Exp3 the Exp strategy run on the estimated losses.

Theorem (Auer, Cesa-Bianchi, Freund and Schapire [2003])
Exp3 satisfies:

$$
R_{n} \leq \sqrt{2 n d \log d}
$$

Moreover for any strategy,

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{p_{t}(i)} \mathbb{1}_{l_{t}=i}
$$

is an unbiased estimate of $\ell_{i, t}$. We call Exp3 the Exp strategy run on the estimated losses.

Theorem (Auer, Cesa-Bianchi, Freund and Schapire [2003])
Exp3 satisfies:

$$
R_{n} \leq \sqrt{2 n d \log d}
$$

Moreover for any strategy,

$$
\sup _{\text {adversaries }} R_{n} \geq \frac{1}{4} \sqrt{n d}+o(\sqrt{n d}) .
$$

## High probability bounds

What about bounds directly on the true regret

$$
\sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{i=1, \ldots, d} \sum_{t=1}^{n} \ell_{i, t} ?
$$

Auer et al. [2003] proposed Exp3.P:

where


Theorem (Auer et al. [2003], Audibert and Bubeck [2011])
Exp3.P satisfies with probability at least $1-\delta$

## High probability bounds

What about bounds directly on the true regret

$$
\sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{i=1, \ldots, d} \sum_{t=1}^{n} \ell_{i, t} ?
$$

Auer et al. [2003] proposed Exp3.P:

$$
p_{t}(i)=(1-\gamma) \frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{j, s}\right)}+\frac{\gamma}{d}
$$

where

Theorem (Auer et al. [2003], Audibert and Bubeck [2011])
Exp3.P satisfies with probability at least $1-\delta$

## High probability bounds

What about bounds directly on the true regret

$$
\sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{i=1, \ldots, d} \sum_{t=1}^{n} \ell_{i, t} ?
$$

Auer et al. [2003] proposed Exp3.P:

$$
p_{t}(i)=(1-\gamma) \frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{j, s}\right)}+\frac{\gamma}{d}
$$

where

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{p_{t}(i)} \mathbb{1}_{l_{t}=i}+\frac{\beta}{p_{t}(i)} .
$$

## Theorem (Auer et al. [2003], Audibert and Bubeck [2011])

Exp3.P satisfies with probability at least $1-\delta$

## High probability bounds

What about bounds directly on the true regret

$$
\sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{i=1, \ldots, d} \sum_{t=1}^{n} \ell_{i, t} ?
$$

Auer et al. [2003] proposed Exp3.P:

$$
p_{t}(i)=(1-\gamma) \frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)}{\sum_{j=1}^{d} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{j, s}\right)}+\frac{\gamma}{d}
$$

where

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{p_{t}(i)} \mathbb{1}_{l_{t}=i}+\frac{\beta}{p_{t}(i)} .
$$

Theorem (Auer et al. [2003], Audibert and Bubeck [2011])
Exp3.P satisfies with probability at least $1-\delta$ :

$$
\sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{i=1, \ldots, d} \sum_{t=1}^{n} \ell_{i, t} \leq 5.15 \sqrt{n d \log \left(d \delta^{-1}\right)}
$$

## Other types of normalization

- INF (Implicitly Normalized Forecaster) is based on a potential function $\psi: \mathbb{R}_{-}^{*} \rightarrow \mathbb{R}_{+}^{*}$ increasing, convex, twice continuously differentiable, and such that $(0,1] \subset \psi\left(\mathbb{R}_{-}^{*}\right)$.
- At each time step INF computes the new probability distribution as follows:
where $C_{t}$ is the unique real number such that $\sum_{i=1}^{d} p_{t}(i)=1$ - $\psi(x)=\exp (n x)+\frac{\gamma}{d}$ corresponds exactly to the Exp3 strategy.


## Theorem (Audibert and Bubeck [2009, 2010])

## Other types of normalization

- INF (Implicitly Normalized Forecaster) is based on a potential function $\psi: \mathbb{R}_{-}^{*} \rightarrow \mathbb{R}_{+}^{*}$ increasing, convex, twice continuously differentiable, and such that $(0,1] \subset \psi\left(\mathbb{R}_{-}^{*}\right)$.
- At each time step INF computes the new probability distribution as follows:

$$
p_{t}(i)=\psi\left(C_{t}-\sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)
$$

where $C_{t}$ is the unique real number such that $\sum_{i=1}^{d} p_{t}(i)=1$.

## Other types of normalization

- INF (Implicitly Normalized Forecaster) is based on a potential function $\psi: \mathbb{R}_{-}^{*} \rightarrow \mathbb{R}_{+}^{*}$ increasing, convex, twice continuously differentiable, and such that $(0,1] \subset \psi\left(\mathbb{R}_{-}^{*}\right)$.
- At each time step INF computes the new probability distribution as follows:

$$
p_{t}(i)=\psi\left(C_{t}-\sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)
$$

where $C_{t}$ is the unique real number such that $\sum_{i=1}^{d} p_{t}(i)=1$.

- $\psi(x)=\exp (\eta x)+\frac{\gamma}{d}$ corresponds exactly to the Exp3 strategy.


## Other types of normalization

- INF (Implicitly Normalized Forecaster) is based on a potential function $\psi: \mathbb{R}_{-}^{*} \rightarrow \mathbb{R}_{+}^{*}$ increasing, convex, twice continuously differentiable, and such that $(0,1] \subset \psi\left(\mathbb{R}_{-}^{*}\right)$.
- At each time step INF computes the new probability distribution as follows:

$$
p_{t}(i)=\psi\left(C_{t}-\sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right),
$$

where $C_{t}$ is the unique real number such that $\sum_{i=1}^{d} p_{t}(i)=1$.

- $\psi(x)=\exp (\eta x)+\frac{\gamma}{d}$ corresponds exactly to the Exp3 strategy.
- $\psi(x)=(-\eta x)^{-1 / 2}+\frac{\gamma}{d}$ is the quadratic INF strategy


## Other types of normalization

- INF (Implicitly Normalized Forecaster) is based on a potential function $\psi: \mathbb{R}_{-}^{*} \rightarrow \mathbb{R}_{+}^{*}$ increasing, convex, twice continuously differentiable, and such that $(0,1] \subset \psi\left(\mathbb{R}_{-}^{*}\right)$.
- At each time step INF computes the new probability distribution as follows:

$$
p_{t}(i)=\psi\left(C_{t}-\sum_{s=1}^{t-1} \tilde{\ell}_{i, s}\right)
$$

where $C_{t}$ is the unique real number such that $\sum_{i=1}^{d} p_{t}(i)=1$.

- $\psi(x)=\exp (\eta x)+\frac{\gamma}{d}$ corresponds exactly to the Exp3 strategy.
- $\psi(x)=(-\eta x)^{-1 / 2}+\frac{\gamma}{d}$ is the quadratic INF strategy


## Theorem (Audibert and Bubeck [2009, 2010])

Quadratic INF satisfies: $R_{n} \leq 2 \sqrt{2 n d}$.

## Extension: contextual bandits

- Contextual bandits: at each time step $t$ one receives a context $s_{t} \in \mathcal{S}$, and one wants to perform as well as the best mapping from contexts to arms:

$$
R_{n}^{\mathcal{S}}=\mathbb{E} \sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{g: \mathcal{S} \rightarrow\{1, \ldots, d\}} \mathbb{E} \sum_{t=1}^{n} \ell_{g\left(s_{t}\right), t}
$$

- A related problem is bandit with experts advice: $N$ experts are playing the game, and the player observes their actions $\xi_{t}^{k}$, $k=1, \ldots, N$. One wants to compete with the best expert:


## Extension: contextual bandits

- Contextual bandits: at each time step $t$ one receives a context $s_{t} \in \mathcal{S}$, and one wants to perform as well as the best mapping from contexts to arms:

$$
R_{n}^{\mathcal{S}}=\mathbb{E} \sum_{t=1}^{n} \ell_{1_{t}, t}-\min _{g: \mathcal{S} \rightarrow\{1, \ldots, d\}} \mathbb{E} \sum_{t=1}^{n} \ell_{g\left(s_{t}\right), t}
$$

- A related problem is bandit with experts advice: $N$ experts are playing the game, and the player observes their actions $\xi_{t}^{k}$, $k=1, \ldots, N$. One wants to compete with the best expert:

$$
R_{n}^{N}=\mathbb{E} \sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{k \in\{1, \ldots, N\}} \mathbb{E} \sum_{t=1}^{n} \ell_{\xi_{t}^{k}, t}
$$

With the bandit feedback $\ell_{l_{t}, t}$ one can build an estimate for

## Extension: contextual bandits

- Contextual bandits: at each time step $t$ one receives a context $s_{t} \in \mathcal{S}$, and one wants to perform as well as the best mapping from contexts to arms:

$$
R_{n}^{\mathcal{S}}=\mathbb{E} \sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{g: \mathcal{S} \rightarrow\{1, \ldots, d\}} \mathbb{E} \sum_{t=1}^{n} \ell_{g\left(s_{t}\right), t} .
$$

- A related problem is bandit with experts advice: $N$ experts are playing the game, and the player observes their actions $\xi_{t}^{k}$, $k=1, \ldots, N$. One wants to compete with the best expert:

$$
R_{n}^{N}=\mathbb{E} \sum_{t=1}^{n} \ell_{l_{t}, t}-\min _{k \in\{1, \ldots, N\}} \mathbb{E} \sum_{t=1}^{n} \ell_{\xi_{t}^{k}, t} .
$$

With the bandit feedback $\ell_{I_{t}, t}$ one can build an estimate for the loss of expert $k$ as $\tilde{\ell}_{t}^{k}=\frac{\ell_{I_{t}, t} \mathbb{1}_{I_{t}=\xi_{t}^{k}}}{p_{t}\left(I_{t}\right)}$. Playing Exp on the set of experts with the above loss estimate yields $R_{n}^{N} \leq \sqrt{2 n d \log N}$.

## Extension: partial monitoring

- Partial monitoring: the received feedback at time $t$ is some signal $S\left(I_{t}, \ell_{t}\right)$, see Cesa-Bianchi and Lugosi [2006].



## Extension: partial monitoring

- Partial monitoring: the received feedback at time $t$ is some signal $S\left(I_{t}, \ell_{t}\right)$, see Cesa-Bianchi and Lugosi [2006].
- A simple interpolation between full info. and bandit feedback is the partial monitoring setting of Mannor and Shamir [2011]: $S\left(I_{t}, \ell_{t}\right)=\left\{\ell_{i, t}, i \in \mathcal{N}\left(I_{t}\right)\right\}$ where $\mathcal{N}:\{1, \ldots, d\} \rightarrow \mathcal{P}(\{1, \ldots, d\})$ is some known neighboorhood mapping.


## Extension: partial monitoring

- Partial monitoring: the received feedback at time $t$ is some signal $S\left(I_{t}, \ell_{t}\right)$, see Cesa-Bianchi and Lugosi [2006].
- A simple interpolation between full info. and bandit feedback is the partial monitoring setting of Mannor and Shamir [2011]: $S\left(I_{t}, \ell_{t}\right)=\left\{\ell_{i, t}, i \in \mathcal{N}\left(I_{t}\right)\right\}$ where $\mathcal{N}:\{1, \ldots, d\} \rightarrow \mathcal{P}(\{1, \ldots, d\})$ is some known neighboorhood mapping. A natural loss estimate in that case is

$$
\tilde{\ell}_{i, t}=\frac{\ell_{i, t} \mathbb{1}_{i \in \mathcal{N}\left(I_{t}\right)}}{\sum_{j \in \mathcal{N}(i)} p_{t}(j)}
$$

Mannor and Shamir [2011] proved that Exp with the above estimate has a regret of order $\sqrt{\alpha n}$ where $\alpha$ is the independence number of the graph associated to $\mathcal{N}$.

Assumption (Robbins [1952])
The sequence of losses $\left(\ell_{t}\right)_{1 \leq t \leq n}$ is a sequence of i.i.d random variables.

For historical reasons in this setting we consider gains rather than losses and we introduce different notation:

- Let $v_{i}$ be the unknown reward distribution underlying arm $i$,
- Let $X_{i, s} \sim \nu_{i}$ be the reward obtained when pulling arm $i$ for

arm $i$ was pulled up to time $t$.
- Thus here


## Assumption (Robbins [1952])

The sequence of losses $\left(\ell_{t}\right)_{1 \leq t \leq n}$ is a sequence of i.i.d random variables.

For historical reasons in this setting we consider gains rather than losses and we introduce different notation:


## Stochastic Assumption

## Assumption (Robbins [1952])

The sequence of losses $\left(\ell_{t}\right)_{1 \leq t \leq n}$ is a sequence of i.i.d random variables.

For historical reasons in this setting we consider gains rather than losses and we introduce different notation:

- Let $\nu_{i}$ be the unknown reward distribution underlying arm $i$, $\mu_{i}$ the mean of $\nu_{i}, \mu^{*}=\max _{1 \leq i \leq d} \mu_{i}$ and $\Delta_{i}=\mu^{*}-\mu_{i}$.

the $s^{\text {th }}$ time, and
arm i was pulled up to time $t$.
- Thus here


## Assumption (Robbins [1952])

The sequence of losses $\left(\ell_{t}\right)_{1 \leq t \leq n}$ is a sequence of i.i.d random variables.

For historical reasons in this setting we consider gains rather than losses and we introduce different notation:

- Let $\nu_{i}$ be the unknown reward distribution underlying arm $i$, $\mu_{i}$ the mean of $\nu_{i}, \mu^{*}=\max _{1 \leq i \leq d} \mu_{i}$ and $\Delta_{i}=\mu^{*}-\mu_{i}$.
- Let $X_{i, s} \sim \nu_{i}$ be the reward obtained when pulling arm $i$ for the $s^{\text {th }}$ time, and $T_{i}(t)=\sum_{s=1}^{t} \mathbb{1}_{I_{s}=i}$ the number of times arm $i$ was pulled up to time $t$.


## Stochastic Assumption

## Assumption (Robbins [1952])

The sequence of losses $\left(\ell_{t}\right)_{1 \leq t \leq n}$ is a sequence of i.i.d random variables.

For historical reasons in this setting we consider gains rather than losses and we introduce different notation:

- Let $\nu_{i}$ be the unknown reward distribution underlying arm $i$, $\mu_{i}$ the mean of $\nu_{i}, \mu^{*}=\max _{1 \leq i \leq d} \mu_{i}$ and $\Delta_{i}=\mu^{*}-\mu_{i}$.
- Let $X_{i, s} \sim \nu_{i}$ be the reward obtained when pulling arm $i$ for the $s^{\text {th }}$ time, and $T_{i}(t)=\sum_{s=1}^{t} \mathbb{1}_{I_{s}=i}$ the number of times arm $i$ was pulled up to time $t$.
- Thus here

$$
R_{n}=n \mu^{*}-\mathbb{E} \sum_{t=1}^{n} \mu_{I_{t}}=\sum_{i=1}^{d} \Delta_{i} \mathbb{E} T_{i}(n)
$$

## Optimism in face of uncertainty

General principle: given some observations from an unknown environment, build (with some probabilistic argument) a set of possible environments $\Omega$, then act as if the real environment was the most favorable one in $\Omega$.

Application to stochastic bandits: given the past rewards, build confidence intervals for the means ( $\mu_{i}$ ) (in particular build upper confidence bounds), then play the arm with the highest upper confidence bound.

## Optimism in face of uncertainty

General principle: given some observations from an unknown environment, build (with some probabilistic argument) a set of possible environments $\Omega$, then act as if the real environment was the most favorable one in $\Omega$.

Application to stochastic bandits: given the past rewards, build confidence intervals for the means ( $\mu_{i}$ ) (in particular build upper confidence bounds), then play the arm with the highest upper confidence bound.

## UCB (Upper Confidence Bounds)

## Theorem (Hoeffding [1963])

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then with probability at least $1-\delta$,

$$
\mathbb{E} X \leq \frac{1}{t} \sum_{s=1}^{t} X_{s}+\sqrt{\frac{\log \delta^{-1}}{2 t}}
$$

This directly suggests the famous UCB strategy of Auer, Cesa-Bianchi and Fischer [2002]

## UCB (Upper Confidence Bounds)

## Theorem (Hoeffding [1963])

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then with probability at least $1-\delta$,

$$
\mathbb{E} X \leq \frac{1}{t} \sum_{s=1}^{t} X_{s}+\sqrt{\frac{\log \delta^{-1}}{2 t}} .
$$

This directly suggests the famous UCB strategy of Auer, Cesa-Bianchi and Fischer [2002]:

$$
I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \frac{1}{T_{i}(t-1)} \sum_{s=1}^{T_{i}(t-1)} X_{i, s}+\sqrt{\frac{2 \log t}{T_{i}(t-1)}} .
$$

## UCB (Upper Confidence Bounds)

## Theorem (Hoeffding [1963])

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then with probability at least $1-\delta$,

$$
\mathbb{E} X \leq \frac{1}{t} \sum_{s=1}^{t} X_{s}+\sqrt{\frac{\log \delta^{-1}}{2 t}}
$$

This directly suggests the famous UCB strategy of Auer, Cesa-Bianchi and Fischer [2002]:

$$
I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \frac{1}{T_{i}(t-1)} \sum_{s=1}^{T_{i}(t-1)} X_{i, s}+\sqrt{\frac{2 \log t}{T_{i}(t-1)}} .
$$

Auer et al. proved the following regret bound:

$$
R_{n} \leq \sum_{i: \Delta_{i}>0} \frac{10 \log n}{\Delta_{i}}
$$

Distribution-dependent lower bound
For any $p, q \in[0,1]$, let

$$
\mathrm{kl}(p, q)=p \log \frac{p}{q}+(1-p) \log \frac{1-p}{1-q} .
$$

## Theorem (Lai and Robbins [1985])

Consider a consistent strategy, i.e. s.t. $\forall a>0$, we have $\mathbb{E} T_{i}(n)=o\left(n^{a}\right)$ if $\Delta_{i}>0$. Then for any Bernoulli reward distributions,

## Note that



## Distribution-dependent lower bound

For any $p, q \in[0,1]$, let

$$
\mathrm{kl}(p, q)=p \log \frac{p}{q}+(1-p) \log \frac{1-p}{1-q} .
$$

## Theorem (Lai and Robbins [1985])

Consider a consistent strategy, i.e. s.t. $\forall a>0$, we have $\mathbb{E} T_{i}(n)=o\left(n^{a}\right)$ if $\Delta_{i}>0$. Then for any Bernoulli reward distributions,

$$
\liminf _{n \rightarrow+\infty} \frac{R_{n}}{\log n} \geq \sum_{i: \Delta_{i}>0} \frac{\Delta_{i}}{\operatorname{kl}\left(\mu_{i}, \mu^{*}\right)}
$$

Note that


## Distribution-dependent lower bound

For any $p, q \in[0,1]$, let

$$
\mathrm{kl}(p, q)=p \log \frac{p}{q}+(1-p) \log \frac{1-p}{1-q}
$$

## Theorem (Lai and Robbins [1985])

Consider a consistent strategy, i.e. s.t. $\forall a>0$, we have $\mathbb{E} T_{i}(n)=o\left(n^{a}\right)$ if $\Delta_{i}>0$. Then for any Bernoulli reward distributions,

$$
\liminf _{n \rightarrow+\infty} \frac{R_{n}}{\log n} \geq \sum_{i: \Delta_{i}>0} \frac{\Delta_{i}}{\operatorname{kl}\left(\mu_{i}, \mu^{*}\right)}
$$

Note that

$$
\frac{1}{2 \Delta_{i}} \geq \frac{\Delta_{i}}{\mathrm{kl}\left(\mu_{i}, \mu^{*}\right)} \geq \frac{\mu^{*}\left(1-\mu^{*}\right)}{2 \Delta_{i}}
$$

## Theorem (Chernoff's inequality)

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then

$$
\mathbb{P}\left(\frac{1}{t} \sum_{s=1}^{t} X_{s} \leq \mathbb{E} X-\epsilon\right) \leq \exp (-t \mathrm{kl}(\mathbb{E} X-\epsilon, \mathbb{E} X))
$$

In particular this implies that with probability at least $1-\delta$ :


## Theorem (Chernoff's inequality)

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then

$$
\mathbb{P}\left(\frac{1}{t} \sum_{s=1}^{t} X_{s} \leq \mathbb{E} X-\epsilon\right) \leq \exp (-t \mathrm{kl}(\mathbb{E} X-\epsilon, \mathbb{E} X))
$$

In particular this implies that with probability at least $1-\delta$ :

$$
\mathbb{E} X \leq \max \left\{q \in[0,1]: \mathrm{kl}\left(\frac{1}{t} \sum_{s=1}^{t} X_{s}, q\right) \leq \frac{\log \delta^{-1}}{t}\right\}
$$

## Theorem (Chernoff's inequality)

Let $X, X_{1}, \ldots, X_{t}$ be i.i.d random variables in $[0,1]$, then

$$
\mathbb{P}\left(\frac{1}{t} \sum_{s=1}^{t} X_{s} \leq \mathbb{E} X-\epsilon\right) \leq \exp (-t \mathrm{kl}(\mathbb{E} X-\epsilon, \mathbb{E} X))
$$

In particular this implies that with probability at least $1-\delta$ :

$$
\mathbb{E} X \leq \max \left\{q \in[0,1]: \mathrm{kl}\left(\frac{1}{t} \sum_{s=1}^{t} X_{s}, q\right) \leq \frac{\log \delta^{-1}}{t}\right\} .
$$



Thus Chernoff's bound suggests the KL-UCB strategy of Garivier and Cappé [2011] (see also Honda and Takemura [2010], Maillard, Munos and Stoltz [2011]) :

$$
\begin{aligned}
I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \max & \{q \in[0,1]: \\
\mathrm{kl} & \left.\left(\frac{1}{T_{i}(t-1)} \sum_{s=1}^{T_{i}(t-1)} X_{i, s}, q\right) \leq \frac{(1+\epsilon) \log t}{T_{i}(t-1)}\right\}
\end{aligned}
$$

Garivier and Cappé proved the following regret bound for $n$ large enough:

Thus Chernoff's bound suggests the KL-UCB strategy of Garivier and Cappé [2011] (see also Honda and Takemura [2010], Maillard, Munos and Stoltz [2011]) :

$$
\begin{aligned}
I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \max & \{q \in[0,1]: \\
\mathrm{kl} & \left.\left(\frac{1}{T_{i}(t-1)} \sum_{s=1}^{T_{i}(t-1)} X_{i, s}, q\right) \leq \frac{(1+\epsilon) \log t}{T_{i}(t-1)}\right\}
\end{aligned}
$$

Garivier and Cappé proved the following regret bound for $n$ large enough:

Thus Chernoff's bound suggests the KL-UCB strategy of Garivier and Cappé [2011] (see also Honda and Takemura [2010], Maillard, Munos and Stoltz [2011]) :

$$
\begin{aligned}
I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \max & \{q \in[0,1]: \\
& \left.\mathrm{kl}\left(\frac{1}{T_{i}(t-1)} \sum_{s=1}^{T_{i}(t-1)} X_{i, s}, q\right) \leq \frac{(1+\epsilon) \log t}{T_{i}(t-1)}\right\} .
\end{aligned}
$$

Garivier and Cappé proved the following regret bound for $n$ large enough:

$$
R_{n} \leq \sum_{i: \Delta_{i}>0}(1+2 \epsilon) \frac{\Delta_{i}}{\mathrm{kl}\left(\mu_{i}, \mu^{*}\right)} \log n .
$$

## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.

The first theoretical guarantee for this strategy was provided in Agrawal and Goyal [2012], and in Kaufmann, Korda, and Munos [2012] it was proved that it attains essentially the same regret than KL-UCB. For the Bayesian regret one can say much more:
$\square$

## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.
- Let $\pi_{i, t}$ be the posterior distribution for $\mu_{i}$ at the $t^{t h}$ round.

The first theoretical guarantee for this strategy was provided in Agrawal and Goyal [2012], and in Kaufmann, Korda, and Munos [2012] it was proved that it attains essentially the same regret than KL-UCB. For the Bayesian regret one can say much more:
$\square$

## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.
- Let $\pi_{i, t}$ be the posterior distribution for $\mu_{i}$ at the $t^{t h}$ round.
- Let $\theta_{i, t} \sim \pi_{i, t}$ (independently from the past given $\pi_{i, t}$ ).

The first theoretical guarantee for this strategy was provided in Agrawal and Goyal [2012], and in Kaufmann, Korda, and Munos [2012] it was proved that it attains essentially the same regret than KL-UCB. For the Bayesian regret one can say much more:


## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.
- Let $\pi_{i, t}$ be the posterior distribution for $\mu_{i}$ at the $t^{t h}$ round.
- Let $\theta_{i, t} \sim \pi_{i, t}$ (independently from the past given $\pi_{i, t}$ ).
- $I_{t} \in \operatorname{argmax}_{i=1, \ldots, d} \theta_{i, t}$.



## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.
- Let $\pi_{i, t}$ be the posterior distribution for $\mu_{i}$ at the $t^{t h}$ round.
- Let $\theta_{i, t} \sim \pi_{i, t}$ (independently from the past given $\pi_{i, t}$ ).
- $I_{t} \in \operatorname{argmax}_{i=1, \ldots, d} \theta_{i, t}$.

The first theoretical guarantee for this strategy was provided in Agrawal and Goyal [2012], and in Kaufmann, Korda, and Munos [2012] it was proved that it attains essentially the same regret than KL-UCB. For the Bayesian regret one can say much more:


## A non-UCB strategy: Thompson's sampling

In Thompson [1933] the following strategy was proposed for the case of Bernoulli distributions:

- Assume a uniform prior on the parameters $\mu_{i} \in[0,1]$.
- Let $\pi_{i, t}$ be the posterior distribution for $\mu_{i}$ at the $t^{t h}$ round.
- Let $\theta_{i, t} \sim \pi_{i, t}$ (independently from the past given $\pi_{i, t}$ ).
- $I_{t} \in \operatorname{argmax}_{i=1, \ldots, d} \theta_{i, t}$.

The first theoretical guarantee for this strategy was provided in Agrawal and Goyal [2012], and in Kaufmann, Korda, and Munos [2012] it was proved that it attains essentially the same regret than KL-UCB. For the Bayesian regret one can say much more:

## Theorem (Russo and van Roy [2013], Bubeck and Liu [2013])

For any prior distribution Thompson Sampling has a Bayesian regret smaller than $14 \sqrt{n K}$.

The standard UCB works for all $\sigma^{2}$ - subgaussian distributions (not only bounded distributions), i.e. such that

$$
\mathbb{E} \exp (\lambda(X-\mathbb{E} X)) \leq \frac{\sigma^{2} \lambda^{2}}{2}, \forall \lambda \in \mathbb{R}
$$

It is easy to see that this is equivalent to

What happens for distributions with heavier tails? Can we get logarithmic regret if the distributions only have a finite variance?

The standard UCB works for all $\sigma^{2}$ - subgaussian distributions (not only bounded distributions), i.e. such that

$$
\mathbb{E} \exp (\lambda(X-\mathbb{E} X)) \leq \frac{\sigma^{2} \lambda^{2}}{2}, \forall \lambda \in \mathbb{R}
$$

It is easy to see that this is equivalent to

$$
\exists \alpha>0 \text { s.t. } \mathbb{E} \exp \left(\alpha X^{2}\right)<+\infty .
$$

What happens for distributions with heavier tails? Can we get logarithmic regret if the distributions only have a finite variance?

The standard UCB works for all $\sigma^{2}$ - subgaussian distributions (not only bounded distributions), i.e. such that

$$
\mathbb{E} \exp (\lambda(X-\mathbb{E} X)) \leq \frac{\sigma^{2} \lambda^{2}}{2}, \forall \lambda \in \mathbb{R}
$$

It is easy to see that this is equivalent to

$$
\exists \alpha>0 \text { s.t. } \mathbb{E} \exp \left(\alpha X^{2}\right)<+\infty .
$$

What happens for distributions with heavier tails? Can we get logarithmic regret if the distributions only have a finite variance?

Median of means, Alon, Gibbons, Matias and Szegedy [2002]

## Lemma

Let $X, X_{1}, \ldots, X_{n}$ be i.i.d random variables such that $\mathbb{E}(X-\mathbb{E} X)^{2} \leq 1$.
Then with probability at least $1-\delta$,

Median of means, Alon, Gibbons, Matias and Szegedy [2002]

## Lemma

Let $X, X_{1}, \ldots, X_{n}$ be i.i.d random variables such that $\mathbb{E}(X-\mathbb{E} X)^{2} \leq 1$.
Then with probability at least $1-\delta$,
$\mathbb{E} X \leq \operatorname{median}\left(\frac{1}{N} \sum_{s=1}^{N} X_{s}, \ldots, \frac{1}{N} \sum_{s=(k-1) N+1}^{k N} X_{s}\right)$

Median of means, Alon, Gibbons, Matias and Szegedy [2002]

## Lemma

Let $X, X_{1}, \ldots, X_{n}$ be i.i.d random variables such that $\mathbb{E}(X-\mathbb{E} X)^{2} \leq 1$. Let $\delta \in(0,1), k=8 \log \delta^{-1}$ and $N=\frac{n}{8 \log \delta^{-1}}$.
Then with probability at least $1-\delta$,

$$
\operatorname{median}\left(\frac{1}{N} \sum_{s=1}^{N} X_{s}, \ldots, \frac{1}{N} \sum_{s=(k-1) N+1}^{k N} x_{s}\right)
$$

Median of means, Alon, Gibbons, Matias and Szegedy [2002]

## Lemma

Let $X, X_{1}, \ldots, X_{n}$ be i.i.d random variables such that $\mathbb{E}(X-\mathbb{E} X)^{2} \leq 1$. Let $\delta \in(0,1), k=8 \log \delta^{-1}$ and $N=\frac{n}{8 \log \delta^{-1}}$.
Then with probability at least $1-\delta$,
$\mathbb{E} X \leq \operatorname{median}\left(\frac{1}{N} \sum_{s=1}^{N} X_{s}, \ldots, \frac{1}{N} \sum_{s=(k-1) N+1}^{k N} X_{s}\right)+8 \sqrt{\frac{8 \log \left(\delta^{-1}\right)}{n}}$.

This suggests a Robust UCB strategy, Bubeck, Cesa-Bianchi and Lugosi [2012]:

$$
\begin{aligned}
& I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \operatorname{median}\left(\frac{1}{N_{i, t}} \sum_{s=1}^{N_{i, t}} X_{i, s}, \ldots, \frac{1}{N_{i, t}} \sum_{s=\left(k_{t}-1\right) N_{i, t}+1}^{k_{t} N_{i, t}} X_{i, s}\right) \\
& \quad+32 \sqrt{\frac{\log t}{T_{i}(t-1)}}
\end{aligned}
$$

with $k_{t}=16 \log t$ and $N_{i, t}=\frac{T_{i}(t-1)}{16 \log t}$.
can be proved for any set of distributions with variance bounded by

This suggests a Robust UCB strategy, Bubeck, Cesa-Bianchi and Lugosi [2012]:

$$
\begin{aligned}
& I_{t} \in \underset{1 \leq i \leq d}{\operatorname{argmax}} \operatorname{median}\left(\frac{1}{N_{i, t}} \sum_{s=1}^{N_{i, t}} X_{i, s}, \ldots, \frac{1}{N_{i, t}} \sum_{s=\left(k_{t}-1\right) N_{i, t}+1}^{k_{t} N_{i, t}} X_{i, s}\right) \\
& \quad+32 \sqrt{\frac{\log t}{T_{i}(t-1)}},
\end{aligned}
$$

with $k_{t}=16 \log t$ and $N_{i, t}=\frac{T_{i}(t-1)}{16 \log t}$. The following regret bound can be proved for any set of distributions with variance bounded by 1 :

$$
R_{n} \leq c \sum_{i: \Delta_{i}>0} \frac{\log n}{\Delta_{i}}
$$

## More extensions

- Slowly changing distributions over time, e.g. Garivier and Moulines (2008).
- Distribution-free regret: UCB has a regret always bounded as $R_{n} \leq c \sqrt{n d \log n}$. Furthermore one can prove that for any strategy there exists a set of distributions such that $R_{n} \geq \frac{1}{20} \sqrt{n d}$. The extraneous logarithmic factor can be removed with MOSS (Audibert and Bubeck (2009))
- If $\mu^{*}$ is known then a constant regret is achievable, Lai and Robbins (1987), Bubeck, Perchet and Rigollet (2013)
- It is possible to design a strategy with simultaneously $R_{n} \leq c \frac{d}{\Lambda} \log ^{2}(n)$ in the stochastic setting, and $R_{n} \leq c \sqrt{d n} \log ^{3}(n)$ in the adversarial setting, Bubeck and Slivkins (2012)
- Bandits with switching cost, Dekel, Ding, Koren and Peres (2013): optimal regret is $\Theta\left(n^{2 / 3}\right)$.


## More extensions

- Slowly changing distributions over time, e.g. Garivier and Moulines (2008).
- Distribution-free regret: UCB has a regret always bounded as $R_{n} \leq c \sqrt{n d \log n}$. Furthermore one can prove that for any strategy there exists a set of distributions such that $R_{n} \geq \frac{1}{20} \sqrt{n d}$. The extraneous logarithmic factor can be removed with MOSS (Audibert and Bubeck (2009)).

Robbins (1987), Bubeck, Perchet and Rigollet (2013)

- It is possible to design a strategy with simultaneously $R_{n} \leq c \frac{d}{\Delta} \log ^{2}(n)$ in the stochastic setting, and $R_{n} \leq c \sqrt{d n} \log ^{3}(n)$ in the adversarial setting, Bubeck and Slivkins (2012)
- Bandits with switching cost, Dekel, Ding, Koren and Peres (2013) optimal regret is $\Theta\left(n^{2 / 3}\right)$.


## More extensions

- Slowly changing distributions over time, e.g. Garivier and Moulines (2008).
- Distribution-free regret: UCB has a regret always bounded as $R_{n} \leq c \sqrt{n d \log n}$. Furthermore one can prove that for any strategy there exists a set of distributions such that $R_{n} \geq \frac{1}{20} \sqrt{n d}$. The extraneous logarithmic factor can be removed with MOSS (Audibert and Bubeck (2009)).
- If $\mu^{*}$ is known then a constant regret is achievable, Lai and Robbins (1987), Bubeck, Perchet and Rigollet (2013).
$R_{n} \leq c \sqrt{d n} \log ^{3}(n)$ in the adversarial setting, Bubeck and Slivkins (2012)
- Bandits with switching cost, Dekel, Ding, Koren and Peres
$\square$


## More extensions

- Slowly changing distributions over time, e.g. Garivier and Moulines (2008).
- Distribution-free regret: UCB has a regret always bounded as $R_{n} \leq c \sqrt{n d \log n}$. Furthermore one can prove that for any strategy there exists a set of distributions such that $R_{n} \geq \frac{1}{20} \sqrt{n d}$. The extraneous logarithmic factor can be removed with MOSS (Audibert and Bubeck (2009)).
- If $\mu^{*}$ is known then a constant regret is achievable, Lai and Robbins (1987), Bubeck, Perchet and Rigollet (2013).
- It is possible to design a strategy with simultaneously $R_{n} \leq c \frac{d}{\Delta} \log ^{2}(n)$ in the stochastic setting, and $R_{n} \leq c \sqrt{d n} \log ^{3}(n)$ in the adversarial setting, Bubeck and Slivkins (2012).
- Bandits with switching cost, Dekel, Ding, Koren and Peres (2013): optimal regret is $\Theta($


## More extensions

- Slowly changing distributions over time, e.g. Garivier and Moulines (2008).
- Distribution-free regret: UCB has a regret always bounded as $R_{n} \leq c \sqrt{n d \log n}$. Furthermore one can prove that for any strategy there exists a set of distributions such that $R_{n} \geq \frac{1}{20} \sqrt{n d}$. The extraneous logarithmic factor can be removed with MOSS (Audibert and Bubeck (2009)).
- If $\mu^{*}$ is known then a constant regret is achievable, Lai and Robbins (1987), Bubeck, Perchet and Rigollet (2013).
- It is possible to design a strategy with simultaneously $R_{n} \leq c \frac{d}{\Delta} \log ^{2}(n)$ in the stochastic setting, and $R_{n} \leq c \sqrt{d n} \log ^{3}(n)$ in the adversarial setting, Bubeck and Slivkins (2012).
- Bandits with switching cost, Dekel, Ding, Koren and Peres (2013): optimal regret is $\Theta\left(n^{2 / 3}\right)$.


## $\mathcal{X}$-armed bandits

Stochastic multi-armed bandit where $\{1, \ldots, K\}$ is replaced by a metric space $\mathcal{X}$. At time $t_{\text {, select }} x_{t} \in \mathcal{X}$, then receive a random variable $Y_{t} \in[0,1]$ such that $\mathbb{E}\left[Y_{t} \mid x_{t}\right]=f\left(x_{t}\right)$.

The regret is defined as:

The standard assumption in this context if that $f$ is Lipschitz.

## $\mathcal{X}$-armed bandits

Stochastic multi-armed bandit where $\{1, \ldots, K\}$ is replaced by a metric space $\mathcal{X}$. At time $t$, select $x_{t} \in \mathcal{X}$, then receive a random variable $Y_{t} \in[0,1]$ such that $\mathbb{E}\left[Y_{t} \mid x_{t}\right]=f\left(x_{t}\right)$.

The regret is defined as:

The standard assumption in this context if that $f$ is Lipschitz.

## $\mathcal{X}$-armed bandits

Stochastic multi-armed bandit where $\{1, \ldots, K\}$ is replaced by a metric space $\mathcal{X}$. At time $t$, select $x_{t} \in \mathcal{X}$, then receive a random variable $Y_{t} \in[0,1]$ such that $\mathbb{E}\left[Y_{t} \mid x_{t}\right]=f\left(x_{t}\right)$.

The regret is defined as:

$$
R_{n}=n \sup _{x \in \mathcal{X}} f(x)-\mathbb{E} \sum_{t=1}^{n} f\left(x_{t}\right)
$$

The standard assumption in this context if that $f$ is Lipschitz.

## $\mathcal{X}$-armed bandits

Stochastic multi-armed bandit where $\{1, \ldots, K\}$ is replaced by a metric space $\mathcal{X}$. At time $t$, select $x_{t} \in \mathcal{X}$, then receive a random variable $Y_{t} \in[0,1]$ such that $\mathbb{E}\left[Y_{t} \mid x_{t}\right]=f\left(x_{t}\right)$.

The regret is defined as:

$$
R_{n}=n \sup _{x \in \mathcal{X}} f(x)-\mathbb{E} \sum_{t=1}^{n} f\left(x_{t}\right)
$$

The standard assumption in this context if that $f$ is Lipschitz.

## $\mathcal{X}$-armed bandits

$\mathcal{X}=[0,1]^{D}, \alpha \geq 0$ and mean-payoff function $f$ locally " $\alpha$-smooth" around (any of) its maximum $x^{*}$ (in finite number):

$$
f\left(x^{*}\right)-f(x)=\Theta\left(\left\|x-x^{*}\right\|^{\alpha}\right) \text { as } x \rightarrow x^{*}
$$

## Theorem

Assume that we run HOO (Bubeck, Munos, Stoltz, Szepesvári, 2008 , 2011) or Zooming algorithm (Kleinberg, Slivkins, Upfal, 2008) using the "metric"

## $\mathcal{X}$-armed bandits

$\mathcal{X}=[0,1]^{D}, \alpha \geq 0$ and mean-payoff function $f$ locally " $\alpha$-smooth" around (any of) its maximum $x^{*}$ (in finite number):

$$
f\left(x^{*}\right)-f(x)=\Theta\left(\left\|x-x^{*}\right\|^{\alpha}\right) \text { as } x \rightarrow x^{*} .
$$

## Theorem

Assume that we run HOO (Bubeck, Munos, Stoltz, Szepesvári, 2008, 2011) or Zooming algorithm (Kleinberg, Slivkins, Upfal, 2008) using the "metric" $\rho(x, y)=\|x-y\|^{\beta}$.

- Known smoothness: independent of the dimension $D$
- Smonthness underestimated:
- Smoothness overestimated: UCT (Kocsis and Szepesvári 2006) corresponds to


## $\mathcal{X}$-armed bandits

$\mathcal{X}=[0,1]^{D}, \alpha \geq 0$ and mean-payoff function $f$ locally " $\alpha$-smooth" around (any of) its maximum $x^{*}$ (in finite number):

$$
f\left(x^{*}\right)-f(x)=\Theta\left(\left\|x-x^{*}\right\|^{\alpha}\right) \text { as } x \rightarrow x^{*}
$$

## Theorem

Assume that we run HOO (Bubeck, Munos, Stoltz, Szepesvári, 2008, 2011) or Zooming algorithm (Kleinberg, Slivkins, Upfal, 2008) using the "metric" $\rho(x, y)=\|x-y\|^{\beta}$.

- Known smoothness: $\beta=\alpha$. $R_{n}=\tilde{O}(\sqrt{n})$, i.e., the rate is independent of the dimension $D$.
- Smoothness underestimated:
- Smoothness overestimated: $\beta>\alpha$. No guarantee. Note: UCT (Kocsis and Szepesvári 2006) corresponds to


## $\mathcal{X}$-armed bandits

$\mathcal{X}=[0,1]^{D}, \alpha \geq 0$ and mean-payoff function $f$ locally " $\alpha$-smooth" around (any of) its maximum $x^{*}$ (in finite number):

$$
f\left(x^{*}\right)-f(x)=\Theta\left(\left\|x-x^{*}\right\|^{\alpha}\right) \text { as } x \rightarrow x^{*}
$$

## Theorem

Assume that we run HOO (Bubeck, Munos, Stoltz, Szepesvári, 2008, 2011) or Zooming algorithm (Kleinberg, Slivkins, Upfal, 2008) using the "metric" $\rho(x, y)=\|x-y\|^{\beta}$.

- Known smoothness: $\beta=\alpha$. $R_{n}=\tilde{O}(\sqrt{n})$, i.e., the rate is independent of the dimension $D$.
- Smoothness underestimated: $\beta<\alpha$.

$$
R_{n}=\tilde{O}\left(n^{(d+1) /(d+2)}\right) \text { where } d=D\left(\frac{1}{\beta}-\frac{1}{\alpha}\right) \text {. }
$$

## $\mathcal{X}$-armed bandits

$\mathcal{X}=[0,1]^{D}, \alpha \geq 0$ and mean-payoff function $f$ locally " $\alpha$-smooth" around (any of) its maximum $x^{*}$ (in finite number):

$$
f\left(x^{*}\right)-f(x)=\Theta\left(\left\|x-x^{*}\right\|^{\alpha}\right) \text { as } x \rightarrow x^{*}
$$

## Theorem

Assume that we run HOO (Bubeck, Munos, Stoltz, Szepesvári, 2008, 2011) or Zooming algorithm (Kleinberg, Slivkins, Upfal, 2008) using the "metric" $\rho(x, y)=\|x-y\|^{\beta}$.

- Known smoothness: $\beta=\alpha$. $R_{n}=\tilde{O}(\sqrt{n})$, i.e., the rate is independent of the dimension $D$.
- Smoothness underestimated: $\beta<\alpha$.

$$
R_{n}=\tilde{O}\left(n^{(d+1) /(d+2)}\right) \text { where } d=D\left(\frac{1}{\beta}-\frac{1}{\alpha}\right) \text {. }
$$

- Smoothness overestimated: $\beta>\alpha$. No guarantee. Note: UCT (Kocsis and Szepesvári 2006) corresponds to $\beta=+\infty$.



## Combinatorial prediction game

Adversary


Player

## Combinatorial prediction game

Adversary


Player $\longrightarrow$


## Combinatorial prediction game



Player $\longrightarrow$


## Combinatorial prediction game



Player $\longrightarrow$


## Combinatorial prediction game



Player $\longrightarrow$

loss suffered: $\ell_{2}+\ell_{7}+\ldots+\ell_{d}$

## Combinatorial prediction game



Player $\longrightarrow$

loss suffered: $\ell_{2}+\ell_{7}+\ldots+\ell_{d}$

## Combinatorial prediction game



Player $\longrightarrow$

loss suffered: $\ell_{2}+\ell_{7}+\ldots+\ell_{d}$

## Combinatorial prediction game



Player $\longrightarrow$

loss suffered: $\ell_{2}+\ell_{7}+\ldots+\ell_{d}$


Notation

$\xrightarrow{\sim} V_{t} \in \mathcal{S}$, loss suffered: $\ell_{t}^{T} V_{t}$

$$
R_{n}=\mathbb{E} \sum_{t=1}^{n} \ell_{t}^{T} V_{t}-\min _{u \in S} \mathbb{E} \sum_{t=1}^{n} \ell_{t}^{T} u
$$


$\leadsto \leadsto V_{t} \in \mathcal{S}$, loss suffered: $\ell_{t}^{T} V_{t}$


$\leftrightarrow \leadsto V_{t} \in \mathcal{S}$, loss suffered: $\ell_{t}^{T} V_{t}$

$$
R_{n}=\mathbb{E} \sum_{t=1}^{n} \ell_{t}^{T} V_{t}-\min _{u \in S} \mathbb{E} \sum_{t=1}^{n} \ell_{t}^{T} u
$$

## Set of concepts $S \subset\{0,1\}^{d}$



Spanning trees
$k$-sized intervals



$$
V_{t} \sim p_{t}, \quad p_{t} \in \Delta(\mathcal{S})
$$

Then, unbiased estimate $\tilde{\ell}_{t}$ of the loss $\ell_{t}$ :

- $\tilde{\ell}_{t}=\ell_{t}$ in the full information game,
- $\tilde{l}_{i, t}=\frac{\ell_{i, t}}{\sum_{V \in S: V_{i}=1} P_{t}(V)} V_{i, t}$ in the semi-bandit game,
- $\tilde{\ell}_{t}=P_{t}^{+} V_{t} V_{t}^{T} \ell_{t}$, with $P_{t}=\mathbb{E}_{V \sim p_{t}}\left(V V^{T}\right)$ in the bandit game.

$$
V_{t} \sim p_{t}, \quad p_{t} \in \Delta(\mathcal{S})
$$

Then, unbiased estimate $\tilde{\ell}_{t}$ of the loss $\ell_{t}$ :

- $\ell_{t}=\ell_{t}$ in the full information game,
- $\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{\sum_{V \in \mathcal{S}: V_{i}=1} p_{t}(V)} V_{i, t}$ in the semi-bandit game,
- $\tilde{\ell}_{+}=P_{t}^{+} V_{+} V_{t}^{\top} \ell_{+}$, with $P_{t}=\mathbb{E}_{V \sim p_{t}}\left(V V^{T}\right)$ in the bandit game.

$$
V_{t} \sim p_{t}, \quad p_{t} \in \Delta(\mathcal{S})
$$

Then, unbiased estimate $\tilde{\ell}_{t}$ of the loss $\ell_{t}$ :

- $\tilde{\ell}_{t}=\ell_{t}$ in the full information game,


$$
V_{t} \sim p_{t}, \quad p_{t} \in \Delta(\mathcal{S})
$$

Then, unbiased estimate $\tilde{\ell}_{t}$ of the loss $\ell_{t}$ :

- $\tilde{\ell}_{t}=\ell_{t}$ in the full information game,
- $\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{\sum_{V \in \mathcal{S}: V_{i}=1} p_{t}(V)} V_{i, t}$ in the semi-bandit game,


$$
V_{t} \sim p_{t}, \quad p_{t} \in \Delta(\mathcal{S})
$$

Then, unbiased estimate $\tilde{\ell}_{t}$ of the loss $\ell_{t}$ :

- $\tilde{\ell}_{t}=\ell_{t}$ in the full information game,
- $\tilde{\ell}_{i, t}=\frac{\ell_{i, t}}{\sum_{V \in \mathcal{S}: V_{i}=1} p_{t}(V)} V_{i, t}$ in the semi-bandit game,
- $\tilde{\ell}_{t}=P_{t}^{+} V_{t} V_{t}^{T} \ell_{t}$, with $P_{t}=\mathbb{E}_{V \sim p_{t}}\left(V V^{T}\right)$ in the bandit game.


## Loss assumptions

Definition ( $L_{\infty}$ )
We say that the adversary satisfies the $L_{\infty}$ assumption: if $\left\|\ell_{t}\right\|_{\infty} \leq 1$ for all $t=1, \ldots, n$.

## Definition $\left(L_{2}\right)$

We say that the adversary satisfies the $L_{2}$ assumption: if $\ell_{t}^{T} v \leq 1$


## Loss assumptions

## Definition ( $L_{\infty}$ )

We say that the adversary satisfies the $L_{\infty}$ assumption: if $\left\|\ell_{t}\right\|_{\infty} \leq 1$ for all $t=1, \ldots, n$.

## Definition $\left(L_{2}\right)$

We say that the adversary satisfies the $L_{2}$ assumption: if $\ell_{t}^{T} v \leq 1$ for all $t=1, \ldots, n$ and $v \in \mathcal{S}$.

Expanded Exponentially weighted average forecaster (Exp2)

$$
p_{t}(v)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} v\right)}{\sum_{u \in \mathcal{S}} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} u\right)}
$$

- In the full information game, against $L_{2}$ adversaries, we have (for some $\eta$ )

$$
R_{n} \leq \sqrt{2 d n},
$$

which is the optimal rate, Dani, Hayes and Kakade [2008]. - Thus against $L_{\infty}$ adversaries we have

$$
R_{n} \leq d^{3 / 2} \sqrt{2 n}
$$

But this is suboptimal, Koolen, Warmuth and Kivinen [2010] - Audibert, Bubeck and Lugosi [2011] showed that, for any $\eta$, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$ adversary such that:


Expanded Exponentially weighted average forecaster
(Exp2)

$$
p_{t}(v)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} v\right)}{\sum_{u \in \mathcal{S}} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} u\right)}
$$

- In the full information game, against $L_{2}$ adversaries, we have (for some $\eta$ )

$$
R_{n} \leq \sqrt{2 d n}
$$

which is the optimal rate, Dani, Hayes and Kakade [2008].

But this is suboptimal, Koolen, Warmuth and Kivinen [2010] - Audibert, Bubeck and Lugosi [2011] showed that, for any $\eta$, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$ adversary such that:

Expanded Exponentially weighted average forecaster
(Exp2)

$$
p_{t}(v)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} v\right)}{\sum_{u \in \mathcal{S}} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} u\right)}
$$

- In the full information game, against $L_{2}$ adversaries, we have (for some $\eta$ )

$$
R_{n} \leq \sqrt{2 d n},
$$

which is the optimal rate, Dani, Hayes and Kakade [2008].

- Thus against $L_{\infty}$ adversaries we have

$$
R_{n} \leq d^{3 / 2} \sqrt{2 n}
$$

But this is suboptimal, Koolen, Warmuth and Kivinen [2010].
there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$ adversary such that:

Expanded Exponentially weighted average forecaster
(Exp2)

$$
p_{t}(v)=\frac{\exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} v\right)}{\sum_{u \in \mathcal{S}} \exp \left(-\eta \sum_{s=1}^{t-1} \tilde{\ell}_{s}^{T} u\right)}
$$

- In the full information game, against $L_{2}$ adversaries, we have (for some $\eta$ )

$$
R_{n} \leq \sqrt{2 d n},
$$

which is the optimal rate, Dani, Hayes and Kakade [2008].

- Thus against $L_{\infty}$ adversaries we have

$$
R_{n} \leq d^{3 / 2} \sqrt{2 n} .
$$

But this is suboptimal, Koolen, Warmuth and Kivinen [2010].

- Audibert, Bubeck and Lugosi [2011] showed that, for any $\eta$, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$ adversary such that:

$$
R_{n} \geq 0.02 \mathrm{~d}^{3 / 2} \sqrt{n} .
$$

## Legendre function

## Definition

Let $\mathcal{D}$ be a convex subset of $\mathbb{R}^{d}$ with nonempty interior $\operatorname{int}(\mathcal{D})$ and boundary $\partial \mathcal{D}$. We call Legendre any function $F: \mathcal{D} \rightarrow \mathbb{R}$ such that

- $F$ is strictly convex and admits continuous first partial derivatives on $\operatorname{int}(\mathcal{D})$,
- For any $u \in \partial \mathcal{D}$, for any $v \in \operatorname{int}(\mathcal{D})$, we have


## Legendre function

## Definition

Let $\mathcal{D}$ be a convex subset of $\mathbb{R}^{d}$ with nonempty interior $\operatorname{int}(\mathcal{D})$ and boundary $\partial \mathcal{D}$. We call Legendre any function $F: \mathcal{D} \rightarrow \mathbb{R}$ such that

- $F$ is strictly convex and admits continuous first partial derivatives on $\operatorname{int}(\mathcal{D})$,
- For any $u \in \partial \mathcal{D}$, for any $v \in \operatorname{int}(\mathcal{D})$, we have


## Legendre function

## Definition

Let $\mathcal{D}$ be a convex subset of $\mathbb{R}^{d}$ with nonempty interior $\operatorname{int}(\mathcal{D})$ and boundary $\partial \mathcal{D}$. We call Legendre any function $F: \mathcal{D} \rightarrow \mathbb{R}$ such that

- $F$ is strictly convex and admits continuous first partial derivatives on $\operatorname{int}(\mathcal{D})$,
- For any $u \in \partial \mathcal{D}$, for any $v \in \operatorname{int}(\mathcal{D})$, we have

$$
\lim _{s \rightarrow 0, s>0}(u-v)^{T} \nabla F((1-s) u+s v)=+\infty .
$$

## Bregman divergence

## Definition

The Bregman divergence $D_{F}: \mathcal{D} \times \operatorname{int}(\mathcal{D})$ associated to a Legendre function $F$ is defined by

$$
D_{F}(u, v)=F(u)-F(v)-(u-v)^{T} \nabla F(v)
$$

## Definition

The Legendre transform of $F$ is defined by

$$
F^{*}(u)=\sup _{x \in \mathcal{D}} x^{T} u-F(x)
$$

Key property for Legendre functions: $\nabla F^{*}=(\nabla F)^{-1}$.

## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$


## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$


## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$


## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$
(1) $w_{t+1}^{\prime} \in \mathcal{D}$ :

$$
w_{t+1}^{\prime}=\nabla F^{*}\left(\nabla F\left(w_{t}\right)-\tilde{\ell}_{t}\right)
$$



## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$
(1) $w_{t+1}^{\prime} \in \mathcal{D}$ :

$$
w_{t+1}^{\prime}=\nabla F^{*}\left(\nabla F\left(w_{t}\right)-\tilde{\ell}_{t}\right)
$$

(2) $w_{t+1} \in \underset{w \in \operatorname{Conv}(\mathcal{S})}{\operatorname{argmin}} D_{F}\left(w, w_{t+1}^{\prime}\right)$


## Online Stochastic Mirror Descent (OSMD)

Parameter: $F$ Legendre on $\mathcal{D} \supset \operatorname{Conv}(\mathcal{S})$
(1) $w_{t+1}^{\prime} \in \mathcal{D}$ :

$$
w_{t+1}^{\prime}=\nabla F^{*}\left(\nabla F\left(w_{t}\right)-\tilde{\ell}_{t}\right)
$$

(2) $w_{t+1} \in \underset{w \in \operatorname{Conv}(\mathcal{S})}{\operatorname{argmin}} D_{F}\left(w, w_{t+1}^{\prime}\right)$
(3) $p_{t+1} \in \Delta(\mathcal{S}): w_{t+1}=\mathbb{E}_{V \sim p_{t+1}} V$

$$
w_{t+1}^{\prime}
$$



## A little bit of advertising 2



吕
S. Bubeck

Theory of Convex Optimization for Machine Learning arXiv:1405.4980

## General regret bound for OSMD

## Theorem

If $F$ admits a Hessian $\nabla^{2} F$ always invertible then,

$$
R_{n} \lesssim \operatorname{diam}_{D_{F}}(\mathcal{S})+\mathbb{E} \sum_{t=1}^{n} \tilde{\ell}_{t}^{T}\left(\nabla^{2} F\left(w_{t}\right)\right)^{-1} \tilde{\ell}_{t}
$$

## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$



## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$



## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$



## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$


(Full Info: Exp
Semi-Bandit=Bandit: Exp3 Auer et al. [2002]

(Full Info: Component Hedge Koolen, Warmuth and Kivinen [2010]

## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$


(Full Info: Exp
Semi-Bandit=Bandit: Exp3 Auer et al. [2002]

(Full Info: Component Hedge Koolen, Warmuth and Kivinen [2010]

Semi-Bandit: MW
Kale, Reyzin and Schapire [2010]

## Different instances of OSMD: LinExp (Entropy Function)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\frac{1}{\eta} \sum_{i=1}^{d} x_{i} \log x_{i}
$$


$\int$ Full Info: Exp
Semi-Bandit=Bandit: Exp3 Auer et al. [2002]

(Full Info: Component Hedge Koolen, Warmuth and Kivinen [2010]

Semi-Bandit: MW
Kale, Reyzin and Schapire [2010]
Bandit: bad algorithm!

Different instances of OSMD: LinINF (Exchangeable Hessian)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\sum_{i=1}^{d} \int_{0}^{x_{i}} \psi^{-1}(s) d s
$$

Different instances of OSMD: LinINF (Exchangeable Hessian)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\sum_{i=1}^{d} \int_{0}^{x_{i}} \psi^{-1}(s) d s
$$



INF, Audibert and Bubeck [2009]

Different instances of OSMD: LinINF (Exchangeable Hessian)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\sum_{i=1}^{d} \int_{0}^{x_{i}} \psi^{-1}(s) d s
$$



INF, Audibert and Bubeck [2009]


$$
\left\{\begin{array}{c}
\psi(x)=\exp (\eta x): \operatorname{LinExp} \\
\psi(x)=(-\eta x)^{-q}, q>1
\end{array}\right.
$$

Different instances of OSMD: LinINF (Exchangeable Hessian)

$$
\mathcal{D}=[0,+\infty)^{d}, F(x)=\sum_{i=1}^{d} \int_{0}^{x_{i}} \psi^{-1}(s) d s
$$



INF, Audibert and Bubeck [2009]


$$
\left\{\begin{array}{l}
\psi(x)=\exp (\eta x): \operatorname{LinExp} \\
\psi(x)=(-\eta x)^{-q}, q>1: \text { LinPoly }
\end{array}\right.
$$

$\mathcal{D}=\operatorname{Conv}(\mathcal{S})$, then

$$
w_{t+1} \in \underset{w \in \mathcal{D}}{\operatorname{argmin}}\left(\sum_{s=1}^{t} \tilde{\ell}_{s}^{T} w+F(w)\right)
$$

## Particularly interesting choice:

Abernethy, Hazan and Rakhlin [2008]

## Different instances of OSMD: Follow the regularized leader

$\mathcal{D}=\operatorname{Conv}(\mathcal{S})$, then

$$
w_{t+1} \in \underset{w \in \mathcal{D}}{\operatorname{argmin}}\left(\sum_{s=1}^{t} \tilde{\ell}_{s}^{T} w+F(w)\right)
$$

Particularly interesting choice: $F$ self-concordant barrier function, Abernethy, Hazan and Rakhlin [2008]

## Theorem (Koolen, Warmuth and Kivinen [2010])

In the full information game, the LinExp strategy (with well-chosen parameters) satisfies for any concept class $S \subset\{0,1\}^{d}$ and any $L_{\infty}$-adversary:

$$
R_{n} \leq d \sqrt{2 n}
$$

Moreover for any strategy, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$-adversary such that:

$$
R_{n} \geq 0.008 d \sqrt{n}
$$

Minimax regret for the semi-bandit game

## Theorem (Audibert, Bubeck and Lugosi [2011])

In the semi-bandit game, the LinExp strategy (with well-chosen parameters) satisfies for any concept class $S \subset\{0,1\}^{d}$ and any $L_{\infty}$-adversary:

$$
R_{n} \leq d \sqrt{2 n}
$$

Moreover for any strategy, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$-adversary such that:

$$
R_{n} \geq 0.008 d \sqrt{n}
$$

Minimax regret for the bandit game

For the bandit game the situation becomes trickier.

- First it appears necessary to add some sort of forced exploration on $S$ to control third order error terms in the regret bound.
- Second, the control of the quadratic term $\tilde{\ell}_{t}^{\top}\left(\nabla^{2} F\left(w_{t}\right)\right)^{-1} \tilde{\ell}_{t}$ is much more involved than previously.


## Minimax regret for the bandit game

For the bandit game the situation becomes trickier.

- First it appears necessary to add some sort of forced exploration on $S$ to control third order error terms in the regret bound.
- Second, the control of the quadratic term is much more involved than previously.


## Minimax regret for the bandit game

For the bandit game the situation becomes trickier.

- First it appears necessary to add some sort of forced exploration on $S$ to control third order error terms in the regret bound.
- Second, the control of the quadratic term $\tilde{\ell}_{t}^{T}\left(\nabla^{2} F\left(w_{t}\right)\right)^{-1} \tilde{\ell}_{t}$ is much more involved than previously.


## John's distribution

## Theorem (John's Theorem)

Let $\mathcal{K} \subset \mathbb{R}^{d}$ be a convex set. If the ellipsoid $\mathcal{E}$ of minimal volume enclosing $\mathcal{K}$ is the unit ball in some norm derived from a scalar product $\langle\cdot, \cdot\rangle$, then there exists $M \leq d(d+1) / 2+1$ contact points $u_{1}, \ldots, u_{M}$ between $\mathcal{E}$ and $\mathcal{K}$, and $\mu \in \Delta_{M}$ (the simplex of dimension $M-1$ ), such that

$$
x=d \sum_{i=1}^{M} \mu_{i}\left\langle x, u_{i}\right\rangle u_{i}, \forall x \in \mathbb{R}^{d}
$$



## Minimax regret for the bandit game

## Theorem (Audibert, Bubeck and Lugosi [2011], Bubeck, <br> Cesa-Bianchi and Kakade [2012])

In the bandit game, the Exp2 strategy with John's exploration satisfies for any concept class $S \subset\{0,1\}^{d}$ and any $L_{\infty}$-adversary:

$$
R_{n} \leq 4 d^{2} \sqrt{n}
$$

and respectively $R_{n} \leq 4 d \sqrt{n}$ for an $L_{2}$-adversary.
Moreover for any strategy, there exists a subset $S \subset\{0,1\}^{d}$ and an $L_{\infty}$-adversary such that:

$$
R_{n} \geq 0.01 d^{3 / 2} \sqrt{n}
$$

For $L_{2}$-adversaries the lower bound is $0.05 \min (n, d \sqrt{n})$.
Conjecture: for an $L_{\infty}$-adversary the correct order of magnitude is $d^{3 / 2} \sqrt{n}$ and it can be attained with OSMD.

